# Catch-at-age data, and a review of natural mortality, for ling 

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Final Research Report for Ministry of Fisheries Research Project MID9801 Objectives 1, 3, 4, \& 5

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Catch-at-age data, and a review of natural mortality, for ling

MID9801
Peter Horn

## 7. Executive Summary

This report describes the final results from Objectives 1, 3, 4, and 5 of project MID9801, relating to ling only. Results from Objective 2 were reported by Harley (1999). New series of catch-at-age data from trawl surveys and commercial longline samples are presented for ling from the Chatham Rise and Southern Plateau. It appears feasible to use the trawl survey samples to develop time series of recruitment indices. A comparison of age-length keys for ling caught by commercial longline and research trawl indicated no differences between fishing methods on the Southern Plateau, or for females on the Chatham Rise, but a clear difference for males from the Chatham Rise. Existing age data were used to review the estimate of natural mortality for ling. Although, no samples were available from unexploited stocks, it seems likely that the currently used value of 0.18 is probably close to the true value.

## 8. Objectives

1. To determine the catch at age from hake fisheries in HAK 1, 4 and 7 and ling fisheries in LIN $3 \& 4,5 \& 6$ and 7 in 1997/98 from samples collected at sea by Scientific Observers and from other sources, with a target coefficient of variation (c.v.) of $30 \%$ for each Fishstock (mean weighted c.v. across all age classes).
2. To update the standardised catch and effort analyses from the ling longline fisheries (LIN $3 \& 4,5 \& 6$ ) with the addition of data up to the end of the 1997/98 fishing year.
3. To determine the feasibility of developing time series of recruitment indices for hake and ling from trawl survey data from Chatham Rise and Southern Plateau.
4. To determine ling age-length keys from trawl surveys and commercial longline fisheries.
5. To review the estimates of natural mortality for hake and ling based on ageing data from the fisheries.
6. To update the stock assessments of hake (HAK 1 and 4) and ling (LIN 3 \& 4 and 5 \& 6) including estimating biomass and yields.

## 9. Methods

1. To determine the catch at age from hake fisheries in HAK 1,4 and 7 and ling fisheries in LIN $3 \& 4,5 \& 6$ and 7 in 1997/98 from samples collected at sea by Scientific Observers and from other sources, with a target coefficient of variation (c.v.) of $30 \%$ for each Fishstock (mean weighted c.v. across all age classes).

As agreed in the "Exceptions and Deviations" section of the tender, no estimates of catch at age from 1997/98 were completed for LIN 5 \& 6 as this had already been completed in 1998 under project MDT9701.

For LIN 3 \& 4, the tender stated a preference to derive catch at age data from commercial longline samples in 1997/98, with a trawl survey sample from January 1999 being an acceptable alternative. No commercial longline otolith samples were available. A sample of 600 otoliths was selected from the survey collection so that the length and sex distribution of the aged fish was proportional to the scaled length frequency from the survey. Otoliths were prepared and read using the method of Horn (1993b). The age data were used to construct age-length keys, which were applied to the scaled length frequency obtained from the survey, to calculate the age structure of the ling population in this area.

For LIN 7, a sample of 600 otoliths was selected from the collection made by Scientific Observers on trawlers off west coast South Island between June and September 1998. The length and sex distribution of the aged fish was proportional to the scaled length frequency from the commercial fishery. Otoliths were prepared and read using the method of Horn (1993b). The age data were used to construct agelength keys, which were applied to the scaled length frequency to calculate the age structure of the ling taken by trawl in this area.
2. To update the standardised catch and effort analyses from the ling longline fisheries (LIN $3 \& 4,5 \& 6$ ) with the addition of data up to the end of the 1997/98 fishing year.

Full details of the analyses of CPUE data from longline fisheries in LIN 3\&4, LIN 5\&6, and LIN 6B are given by Harley (1999).

## 3. To determine the feasibility of developing time series of recruitment indices for hake and ling from trawl survey data from Chatham Rise and Southern Plateau.

Two trawl survey series which monitor stocks of ling on Chatham Rise and Campbell Plateau are currently being maintained, i.e., a January Chatham Rise survey conducted annually since 1992, and an April-May Campbell Plateau survey conducted four times
since 1992. Catch at age data were already available from some of these surveys (see Horn 1997), and they were derived for the remaining surveys.

Otoliths from three Southern Plateau surveys (TAN9204, TAN9304, TAN9605) and four Chatham Rise surveys (TAN9212, TAN9401, TAN9501, TAN9601) were selected for ageing. Samples of 600 otoliths were selected from each survey sample so that the length and sex distribution of the aged fish was proportional to the scaled length frequency from the survey. Otoliths were prepared and read using the method of Horn (1993b). The age data were used to construct age-length keys, which were applied to the scaled length frequencies obtained from the surveys, to calculate age structures of the ling populations in these areas.

These series will be incorporated into stock models, and used to determine series of recruitment indices for ling from the two areas.

## 4. To determine ling age-length keys from trawl surveys and commercial longline fisheries.

To enable comparisons of commercial longline and research trawl age-length keys from the Chatham Rise and Campbell Plateau, two further samples were prepared and read. Two samples each of 550 otoliths were selected from commercial longline trips in LIN 5\&6 in April-May 1998, and in LIN 3\&4 in January 1999. The otoliths were selected so that the length and sex distribution of the aged fish was proportional to the scaled length frequency from the trip. The age data were used to construct age-length keys, which were applied to the scaled length frequencies, to calculate age structures. Both the sampled commercial trips occurred at the same times as trawl surveys of the respective areas (i.e., TAN9805, TAN9901).

Mean lengths at age were calculated separately by sex for ling from the two commercial longline trips and the two research trawl surveys conducted at the same time. Means were calculated only when there was a minimum of three fish at a particular age. Pairs of mean lengths at age from the same area and sex, but different fishing methods, were compared using $t$-tests.

## 5. To review the estimates of natural mortality for hake and ling based on ageing data from the fisheries.

Existing age data were used to review the estimation of natural mortality for ling. Estimates of $M$ were derived using the methods of Hoenig (1983) and Chapman \& Robson (1960), and from the slope of the right hand limb of the catch curve (Ricker 1975). Full details of the methods used are presented in Appendix B below.

## 6. To update the stock assessments of hake (HAK 1 and 4) and ling (LIN 3 \& 4 and 5 \& 6) including estimating biomass and yields.

Ling stocks LIN $3 \& 4$ and LIN 5\&6 were modelled using the least squares and singlestock MIAEL estimation techniques of Cordue (1993, 1995, 1998). Both these stocks support trawl and longline fisheries, and it is apparent that these two methods have markedly different fishing selectivity ogives. The Middle Depth Species Fisheries Assessment Working Group considered that the use of a single selectivity ogive for a dual fishery was inappropriate. Subsequently, the MIAEL estimation procedure was
modified so that catch histories and selectivities attributable to different fishing methods could be used separately. Modelling results will be presented later.

## 10. Results

1. To determine the catch at age from hake fisheries in HAK 1, 4 and 7 and ling fisheries in LIN 3 \& 4, 5 \& 6 and 7 in 1997/98 from samples collected at sea by Scientific Observers and from other sources, with a target coefficient of variation (c.v.) of $\mathbf{3 0} \%$ for each Fishstock (mean weighted c.v. across all age classes).

The catch at age distribution for ling form the 1998 trawl survey of the Southern Plateau (TAN9805) is given in Appendix A (Table A1). The distribution from the 1999 Chatham Rise trawl survey (TAN9901) is listed in Appendix A (Table A2). The catch at age distribution form the 1997/98 trawl fishery in LIN 7 was presented by Horn \& Ballara (1999).

All the catch at age samples had coefficients of variation of less than $30 \%$.
2. To update the standardised catch and effort analyses from the ling longline fisheries (LIN $3 \& 4,5 \& 6$ ) with the addition of data up to the end of the 1997/98 fishing year.

Full details of the analyses of CPUE data from longline fisheries in LIN 3\&4, LIN 5\&6, and LIN 6B are given by Harley (1999).
3. To determine the feasibility of developing time series of recruitment indices for hake and ling from trawl survey data from Chatham Rise and Southern Plateau.

Appendix A contains catch at age data for ling from the series of four trawl surveys of the Campbell Plateau in autumn 1992, 1993, 1996, and 1998 (Table A1), and from the series of eight trawl surveys of the Chatham Rise conducted annually, in January, from 1992 to 1999 (Table A2).

These series will be incorporated into stock assessment models, and used to determine series of recruitment indices for ling from the two areas. (The series will be reported later with the stock assessment results.)
4. To determine ling age-length keys from trawl surveys and commercial longline fisheries.

Catch at age data for ling from two commercial longline trips (in LIN 5\&6 in AprilMay 1998, and in LIN $3 \& 4$ in January 1999) are compared with data from trawl surveys conducted in the same areas and times (Appendix A, Table A3). It is apparent that, in both areas, the longline fishery catches a greater proportion of older fish than the trawl surveys (Figure A1).

Mean lengths at age were calculated separately by sex for ling from the two commercial longline trips and the two research trawl surveys conducted at the same time. Means were calculated only when there was a minimum of three fish at a
particular age, and are shown in Appendix A (Figure A2). Pairs of mean lengths at age from the same area and sex, but different fishing methods, were compared using $t$ tests. There were no significantly different pairings for Campbell Plateau fish, or for females on the Chatham Rise. However, male ling on the Chatham Rise caught by longline appear to be consistently larger at a particular age than trawl-caught fish. Four individual pairings of means were statistically significant (i.e., ages $10,11,16$, and 17), although for 15 of the 16 ages that could be compared, the mean length was smaller for the trawl survey fish

Age-length keys calculated for ling on the Campbell Plateau, and for female ling on the Chatham Rise, can probably be applied to length-frequency data from either the trawl or longline fisheries. For male ling from the Chatham Rise, the age-length keys appear to be fishery dependent.

## 5. To review the estimates of natural mortality for hake and ling based on ageing data from the fisheries.

Full details of a revision of the estimates of natural mortality for ling are presented in Appendix B below. The current study indicates that the estimate of $M$ used in previous assessments of ling (i.e., 0.18 for both sexes) was probably reasonable. No suitable otolith samples are available from an unexploited ling stock, so it is difficult to confidently update the estimate of $M$. It is suggested that 0.16 is the best estimate of $M$ for females. Given that the $M$ for male ling should probably be slightly higher than that for females, a value of 0.18 is suggested here.

The Middle Depth Species Fisheries Assessment Working Group chose to retain the estimate of $M$ used in previous assessments, i.e., 0.18 for both sexes.

## 11. Publications

Harley, S.J. 1999: Catch per unit effort (CPUE) analysis of the Chatham Rise, Southern Plateau, and Bounty Platform ling (Genypterus blacodes) longline fisheries. N.Z. Fisheries Assessment Research Document 99/31. 26 p.

## 12. Data Storage

All age data are stored on the age database at NIWA, Greta Point, Wellington.

## 13. References

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Horn, P.L. \& Ballara, S.L. 1999: Analysis of longline CPUE, and stock assessment of ling (Genypterus blacodes) off the northwest coast of the South Island (Fishstock LIN 7). N.Z. Fisheries Assessment Research Document 99/27. 31 p.
Pauly, D. 1980: On the interrelationships between natural mortality, growth parameters, and mean environmental temperature in 175 fish stocks. Journal du Conseil International pour l'Exploration de la Mer 39: 175-192.
Ricker, W.E. 1975: Computation and interpretation of biological statistics of fish populations. Bulletin of the Fisheries Research Board, Canada 191.382 p.

## APPENDIX A

## Calculated catch-at-age data

Table A1: Calculated catch at age, by sex, for ling from a series of comparable trawl surveys in the Southern Plateau. Numbers have been scaled to total population in the survey area. Summary statistics show the total number of fish, by sex, that were measured (Meas.mal, Meas.fem) or successfully aged (Aged.mal, Aged.fem), the number of trawl shots that were sampled (Samp.tows), and the mean weighted c.v. across all age classes for the catch at age data (Mean CV)


Table A1: (continued)


Table A2: Calculated catch at age, by sex, for ling from a series of comparable trawl surveys on the Chatham Rise. Numbers have been scaled to total population in the survey area. Summary statistics show the total number of fish, by sex, that were measured (Meas.mal, Meas.fem) or successfully aged (Aged.mal, Aged.fem), the number of trawl shots that were sampled (Samp.tows), and the mean weighted c.v. across all age classes for the catch at age data (Mean CV)
TAN9106
Number $\quad c v$

Table A2: (continued)


Table A3: Calculated catch-at-age, by sex, for comparable samples of ling caught by commercial longline and research trawl in LIN 5\&6 (April-May 1998) and LIN 3\&4 (January 1999). Numbers have been scaled to total population in the survey area (for trawl surveys), or to total sampled catch (for commercial longline). Summary statistics show the total number of fish, by sex, that were measured (Meas.mal, Meas.fem) or successfully aged (Aged.mal, Aged.fem), the number of longline sets or trawl shots that were sampled (Samp.sets/tows), and the mean weighted $c . v$. across all age classes for the catch at age data (Mean CV)

|  | Age | LIN 3 \& 4 |  |  |  | LIN 5 \& 6 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Longline |  | TAN9901 |  | Longline |  | TAN9805 |  |
|  |  | Number | cv | Number | cv | Number | cv | Number | cV |
| Male | 2 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 | 64324 | 0.522 |
|  | 3 | 0 | 0.000 | 268651 | 0.152 | 0 | 0.000 | 270431 | 0.271 |
|  | 4 | 0 | 0.000 | 318130 | 0.145 | 12 | 0.695 | 722914 | 0.185 |
|  | 5 | 4 | 2.286 | 302895 | 0.166 | 112 | . 0.259 | 746835 | 0.186 |
|  | 6 | 16 | 0.713 | 118562 | 0.336 | 135 | 0.229 | 488161 | 0.231 |
|  | 7 | 39 | 0.612 | 101855 | 0.327 | 174 | 0.277 | 608391 | 0.222 |
|  | 8 | 147 | 0.276 | 137953 | 0.252 | 161 | 0.403 | 620412 | 0.220 |
|  | 9 | 231 | 0.229 | 112433 | 0.321 | 84 | 0.512 | 412598 | 0.277 |
|  | 10 | 161 | 0.254 | 103482 | 0.320 | 0 | 0.000 | 259363 | 0.338 |
|  | 11 | 187 | 0.227 | 136852 | 0.267 | 21 | 0.951 | 446643 | 0.268 |
|  | 12 | 210 | 0.225 | 84922 | 0.330 | 77 | 0.583 | 350274 | 0.285 |
|  | 13 | 123 | 0.281 | 64229 | 0.358 | 90 | 0.266 | 85830 | 0.521 |
|  | 14 | 80 | 0.331 | 73096 | 0.403 | 48 | 0.660 | 198515 | 0.358 |
|  | 15 | 67 | 0.430 | 53230 | 0.533 | 170 | 0.278 | 232012 | 0.328 |
|  | 16 | 53 | 0.480 | 93711 | 0.323 | 186 | 0.222 | 241599 | 0.305 |
|  | 17 | 23 | 0.564 | 51421 | 0.357 | 147 | 0.278 | 206884 | 0.370 |
|  | 18 | 16 | 0.642 | 28447 | 0.420 | 45 | 0.771 | 83027 | 0.547 |
| , | 19 | 45 | 0.520 | 27277 | 0.382 | 59 | 0.564 | 55350 | 0.606 |
|  | 20 | 10 | 1.076 | 2209 | 3.350 | 183 | 0.255 | 60647 | 0.686 |
|  | 21 | 30 | 0.652 | 23795 | 0.642 | 96 | 0.415 | 38106 | 0.753 |
|  | 22 | 37 | 0.444 | 56424 | 0.643 | 0 | 0.000 | 16151 | 0.897 |
|  | 23 | 13 | 0.782 | 14405 | 0.612 | 0 | 0.000 | 8814 | 1.487 |
|  | 24 | 27 | 0.706 | 15094 | 0.726 | 102 | 0.296 | 18478 | 1.065 |
|  | 25 | 28 | 0.606 | 9187 | 0.587 | 18 | 0.878 | 0 | 0.000 |
|  | 26 | 13 | 0.931 | 3446 | 0.849 | 29 | 0.728 | 17412 | 1.036 |
|  | 27 | 15 | 0.930 | 7103 | 1.277 | 0 | 0.000 | 0 | 0.000 |
|  | 28 | 22 | 0.689 | 1321 | 1.531 | 22 | 0.739 | 7377 | 1.156 |
|  | 29 | 0 | 0.000 | 0 | 0.000 | 27 | 0.916 | 0 | 0.000 |
|  | 30 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 |
|  | 31 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 |
|  | 32 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 |
|  | 33 | 3 | 2.054 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 |
|  | 34 | 0 | 0.000 | 2358 | 0.582 | 0 | 0.000 | 0 | 0.000 |
|  | 35 | 7 | 1.680 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 |

Table A3: (continued)

| Age | LIN 3 \& 4 |  |  |  | LIN 5 \& 6 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Longline |  | TAN9901 |  | Longline |  | TAN9805 |  |
|  | Number | cv | Number | cv | Number | cV | Number | cv |
| Female 2 | 0 | 0.000 | 7333 | 0.966 | 0 | 0.000 | 98804 | 0.631 |
| 3 | 0 | 0.000 | 219770 | 0.151 | 0 | 0.000 | 439282 | 0.262 |
| 4 | 0 | 0.000 | 304700 | 0.178 | 5 | 1.221 | 634221 | 0.215 |
| 5 | 22 | 0.598 | 163194 | 0.198 | 81 | 0.376 | 769880 | 0.179 |
| 6 | 80 | 0.296 | 126716 | 0.236 | 301 | 0.231 | 525512 | 0.215 |
| 7 | 176 | 0.230 | 133639 | 0.238 | 215 | 0.330 | 525028 | 0.218 |
| 8 | 530 | 0.180 | 138714 | 0.219 | 500 | 0.186 | 508310 | 0.232 |
| 9 | 856 | 0.139 | 106371 | 0.269 | 479 | 0.222 | 601825 | 0.226 |
| 10 | 919 | 0.137 | 136134 | 0.227 | 700 | 0.190 | 376171 | 0.278 |
| 11 | 959 | 0.130 | 97249 | 0.255 | 676 | 0.192 | 322754 | 0.297 |
| 12 | 689 | 0.153 | 75874 | 0.301 | 817 | 0.178 | 321985 | 0.312 |
| 13 | 674 | 0.146 | 49682 | 0.355 | 551 | 0.204 | 299375 | 0.298 |
| 14 | 484 | 0.173 | 30444 | 0.433 | 791 | 0.178 | 229846 | 0.337 |
| 15 | 420 | 0.181 | 26046 | 0.429 | 1281 | 0.138 | 228994 | 0.323 |
| 16 | 190 | 0.260 | 19015 | 0.579 | 687 | 0.184 | 164562 | 0.357 |
| 17 | 132 | 0.304 | 17375 | 0.427 | 618 | 0.200 | 207940 | 0.330 |
| 18 | 153 | 0.276 | 21860 | 0.460 | 569 | 0.208 | 117701 | 0.464 |
| 19 | 67 | 0.434 | 16218 | 1.522 | 373 | 0.251 | 48843 | 0.762 |
| 20 | 71 | 0.422 | 16171 | 0.697 | 230 | 0.289 | 20849 | 1.216 |
| 21 | 25 | 0.573 | 1588 | 1.209 | 391 | 0.230 | 48293 | 0.935 |
| 22 | 70 | 0.410 | 5516 | 0.764 | 122 | 0.388 | 0 | 0.000 |
| 23 | 43 | 0.536 | 1496 | 1.128 | 93 | 0.522 | 12747 | 1.396 |
| 24 | 21 | 0.710 | 2619 | 1.734 | 34 | 0.966 | 12585 | 1.347 |
| 25 | 23 | 0.780 | 9185 | 0.625 | 22 | 0.924 | 0 | 0.000 |
| 26 | 11 | 0.856 | 9773 | 0.690 | 87 | 0.410 | 0 | 0.000 |
| 27 | 28 | 0.598 | 0 | 0.000 | 71 | 0.381 | 0 | 0.000 |
| 28 | 11 | 0.856 | 3666 | 1.230 | 0 | 0.000 | 0 | 0.000 |
| 29 | 5 | 1.519 | 0 | 0.000 | 48 | 0.756 | 0 | 0.000 |
| 30 | 9 | 1.118 | 3030 | 6.927 | 52 | 0.497 | 0 | 0.000 |
| 31 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 |
| 32 | 7 | 1.167 | 0 | 0.000 | 0 | 0.000 | 0 | 0.000 |
| 33 | 7 | 1.101 | 2711 | 0.964 | 0 | 0.000 | 0 | 0.000 |
| Meas.mal | 410 |  | 1079 |  | 608 |  | 816 |  |
| Meas.fem | 1590 |  | 851 |  | 2763 |  | 767 |  |
| Aged.mal | 117 |  | 333 |  | 82 |  | 271 |  |
| Aged.fem | 416 |  | 317 |  | 404 |  | 296 |  |
| Smp.sets/tows | 69 |  | 113 |  | 34 |  | 64 |  |
| Mean CV | 22.3 |  | 28.2 |  | 24.1 |  | 27.8 |  |

LIN 3\&4-longline


LIN 3\&4-traw I survey


LIN 5\&6-Iongline


LIN 5\& 6-trawlsurvey


Figure A1: Calculated proportion-at-age, by sex, for comparable samples of ling caught by commercial longline and research trawl in LIN 5\&6 (April-May 1998) and LIN 3\&4 (January 1999)

Chatham - male


Figure A2: Mean lengths at age, separately by sex, for ling caught by commercial longline (open circles) and by research trawl (closed diamonds), on the Chatham Rise in January 1999 and the Campbell Plateau in April-May 1998. Means were calculated only where $n \geq 3$. Bars show range of size at age

## APPENDIX B

## A review of the estimate of natural mortality for ling

Objective 5 of Project MID9801 required the estimate of natural mortality ( $M$ ) for ling to be reviewed, and revised if possible. A report of this investigation follows, and concludes that the best estimates of $M$ for ling are 0.18 for males and 0.16 for females.

## 1. Methods

Numerous sets of age data, represented as catch-at-age distributions, were available for ling from various areas and years. They comprise samples from trawl surveys of the Chatham Rise from 1989 to 1999, trawl surveys of the Southern Plateau from 1989 to 1998, and the commercial trawl catch off the west coast of the South Island (WCSI) in 1991, 1997, and 1998. Data were also available from two samples from the commercial longline catch on the Southern Plateau (1998) and the Chatham Rise (1999).

Available sets of catch-at-age data were collated by biological stock (as defined by Horn \& Cordue 1996), sex, and sampling method, but split into groups of samples collected before or after February 1993. Numbers at age of all samples within each group were then summed across each age class to produce a single distribution for each sex in each stock, by sampling method. Where several years of data are combined in this way it has the effect of smoothing the data and reducing the influence of any particularly weak or strong year classes.

Estimates of instantaneous natural mortality $(M)$ were derived using the three following methods.

1. $M=-\frac{\log _{e}(p)}{A}$
where $p$ is the proportion of the population that reaches age $A$ or older (Hoenig 1983). Values of $p$ of 0.01 and 0.05 were used here. In an unexploited population, a $p$ of 0.01 is usually used. This method assumes that all age classes in the population are fully vulnerable to the sampling technique. This assumption does not hold for ling; in all samples, some of the younger age classes are not fully recruited. To correct for this, an age at recruitment $R$ was chosen, the proportion $p$ of the fully recruited population (i.e., aged $R$ or older) that reaches age $A$ or older is calculated, and the denominator in Hoenig's equation is replace by $A-R+1$. This is subsequently referred to as the $\mathrm{A}_{\text {max }}$ estimator.
2. $M=\log _{\cdot}\left(\frac{1+a-(1 / n)}{a}\right)$
where $a$ is the mean age above recruitment age and $n$ is the sample size (Chapman \& Robson 1960). For this estimator, age at recruitment ( $R$ ) should be the age at which $100 \%$ of fish are vulnerable to the sampling method (rather than the often used age at $50 \%$ recruitment). This is subsequently
referred to as the CR estimator. A $95 \%$ confidence interval around this estimator is $\pm 2^{*} \sqrt{ }$ var, where var $=\left(1-e^{-\mathrm{CR}}\right)^{2} /\left(n e^{-\mathrm{CR}}\right)$.
3. $\quad M$ is minus the slope of the right hand limb (i.e., points where age is $R$ or older) of the relationship between age and the natural logarithm of the frequency of fish in that age class (Ricker 1975). The regression model defined as R1 by Dunn et al. (1999) was used here, i.e., reject all fish of ages greater than $i_{\max }$, where $i_{\max }$ is the greatest age for which $N_{i} \geq 1$ for all $i \leq i_{\max }$. This is subsequently referred to as the R1 estimator. A 95\% confidence interval around this estimator was taken as $\pm 2 * \mathrm{SE}$ of the slope.

All three methods estimate instantaneous total mortality $(Z)$, rather than $M$. However, if it can be assumed for any particular sample that exploitation (i.e., instantaneous fishing mortality, $F$ ) has been negligible, then $Z$ will approximate $M$.

All three estimators require the initial determination of an age at full recruitment $(R)$. It was not possible to use a single, consistent age at full recruitment for all samples from all areas, owing to differences in sampling methods (i.e., research trawl with 60 mm mesh codend, commercial trawl with 100 mm mesh codend, and commercial longline). However, for valid comparisons between estimates from an individual sample, it is important to use the same age at recruitment in all estimators.

## 2. Results

Details of the sets of samples used in this analysis, and the chosen ages at recruitment $(R)$, are given in Table B1. There was considerable uncertainty about the age at recruitment for the fish from the Chatham Rise (LIN 3 \& 4) and the Southern Plateau (LIN 5 \& 6). Some samples indicated that ling as young as 4 or 5 years old may be fully recruited. However, because the results of all estimators are sensitive to the chosen $R$, it was considered prudent to select higher values, thereby increasing the likelihood that only fully recruited year classes are used in the estimations. The values chosen here vary between stocks, but were the lowest age that appeared to be fully recruited in all samples from a particular stock.

Catch curve plots, and calculated regression lines of the right hand limbs of these distributions, are presented in Figures B1, B2, B3, and B4: Estimates of $Z$ using the three estimators are summarised in Table B1. The estimates have a reasonably wide range, both within and between samples and estimation methods.

For LIN 3 \& 4 (Chatham Rise), there were relatively small differences in trawl survey estimates of $Z$ from the earlier and later periods of exploitation, but any differences did indicate that the earlier period had the greater rate of total mortality. The catch curves (Figure B1) also exhibit only minor difference in shape between periods. There is some indication that the trawl survey catch curves could be better represented by 'broken stick' regressions, with points of discontinuity at about 15 years in the earlier period and 20 years in the later period (Figure B5). Estimates of $Z$ from the single Chatham Rise commercial longline sample (Figure B2) are higher for females than those derived from the trawl surveys, but lower for males. However, the longline fishery may not comprehensively sample males; females comprised more than $70 \%$ of the catch.

Estimates of $Z$ for LIN 5 \& 6 (Southern Plateau) derived from research trawl surveys also exhibit relatively small differences between comparable estimates from the earlier and later periods of exploitation. The catch curves from the two periods are also similar (Figure B3). As in the case of Chatham Rise fish, the trawl survey catch curves from the Southern Plateau could be represented by 'broken stick' regressions, particularly for females. The estimates of $Z$ from the single 1998 commercial longline sample (Figure B2) are consistently lower than those from the research samples taken from 1993-98. As on the Chatham Rise, the longline fishery catches relatively few males.

The commercial trawl samples from LIN 7 (Figure B4) produce estimates of $Z$ which are lower in 1991 than in the latter part of the 1990s.

Comparisons between sexes of estimates of $Z$ derived from the same data set and estimation method exhibit some trends. For all the LIN 7 samples, there is a consistent trend for $Z$ to be greater for males, relative to females. The LIN $3 \& 4$ research samples exhibit a similar trend; however, from research surveys of LIN $5 \& 6$, comparable estimates are generally higher for females. Both commercial longline samples produce estimates of $Z$ which are generally higher for females (although males in both areas are poorly sampled by this method).

## 3. Discussion

Dunn et al. (1999) compared the merits of estimating $M$ using the CR estimator and various regression methods (including R1). They concluded that CR was the most accurate estimator, with R1 being the best of the regression estimators. They did not examine any $A_{\max }$ estimators. The values of variance reported here for the CR and R1 estimators are likely to be underestimates because their calculation assumes no ageing error, constant recruitment, a constant $M$, and little sampling variability. None of these assumptions are likely to be true.

Clearly, estimates derived using all three methods can vary (sometimes quite markedly) given uncertainties in some parameters. The CR estimator is sensitive to the chosen age at recruitment. The R1 estimate can be influenced by where the right limb of the catch curve is started (i.e., the chosen age at recruitment), or by a single outlying point in the data series. To demonstrate the dependence of $Z$ on the selected age at recruitment, results using the CR and R1 estimation methods are presented for several data sets (Figure B6). Both estimators should become asymptotic, assuming that $M$ and $F$ are constant for all fish older than the age of recruitment (A. Dunn, NIWA, pers. comm.). The shapes of the curves in Figure B6 vary considerably, and few exhibit a clear asymptote. If there is a trend in curve shape it is for an initial increase in $Z$, followed by a period where the curve flattens (or even declines), then another period of increase (sometimes followed by a relatively steep decline). The data presented in Figures B5 and B6 raise the possibility that $M$ and/or $F$ are not constant after recruitment, or have changed markedly during the period covered by the samples.

There is some indication from the Chatham Rise trawl samples in Figure B6 that the CR and R1 estimates of $Z$ are approaching asymptotes between about 0.16 and 0.22 over the sections where age at recruitment is assumed to be in the range of 6 to 13 years. This characteristic is particularly apparent for females. The CR estimates are lower than those derived using the R1 estimator, and indicate an asymptote at about 0.18 .

It was considered possible that the 'broken stick' nature of the catch curve regressions for the Chatham Rise samples could be a result of heavy exploitation in 1976-77 when about $45,000 \mathrm{t}$ of ling was believed to have been taken from this region (Horn 1997) primarily by foreign longliners (Annala et al. 1998). However, assuming that fish from about age 9 and older were heavily exploited by this method (as indicated for recent longlining, see Figure B2), then the 'break' in the regression should occur at about age 23 in the 1989-93 plot, and about age 28 in the 1994-99 plot. From Figure B5, it is apparent that the breaks occur somewhere in the range of $14-18$ years in 1989-93, and 19-22 years in 1994-99. For the fishery to be responsible for the apparent change in $Z$, age classes as young as 1-3 years would have to have been heavily exploited by longline operations. This is considered unlikely. As the position of the break shifts by about 5 years between the two samples, and the mean difference between the two sampling periods is also 5 years, it is likely that the change in $Z$ at this point is a true characteristic of the population, rather than an effect related to fishing selectivity.

The regression lines in Figure B5 for fish aged from 5 to about $15-20$ years indicate a $Z$ of about 0.06 for males and 0.15 for females. These values (particularly that for male ling) are considered to be too low. There are numerous possible explanations for these relatively flat curves. Two of the more likely ones are that the younger fish in the range are not fully sampled by the trawl gear, or that there are some relatively strong year classes in the latter part of the range.

The catch curves from trawl surveys of the Southern Plateau could also be fitted using a 'broken stick' regression, with the break somewhere in the range of 14 to 18 years. It is unlikely that heavy exploitation prior to the 1990s could be responsible for this effect; estimated annual landings were generally in the order of 2000 to $3000 t$, with a peak in 1979 of about 6400 t (Horn 1997).

All estimates of $Z$ based on the samples examined here are likely to comprise a nontrivial component of $F$. However, samples from LIN 5, 6, and 7 collected before 1993 should be the least influenced in this regard because landings from these areas were relatively low (Annala et al. 1998). Estimates of $Z$ from these samples using the CR and R1 methods range from 0.20 to 0.34 . Using the $\mathrm{A}_{\max }$ method and a $p$ of 0.05 gave a range of 0.21 to 0.30 . While the $\mathrm{A}_{\max }$ estimate may be a better estimate of $M$ than those from CR and R1 because it does account for some component of fishing mortality, the true extent of $F$ is unknown, so the chosen $p$ may be inappropriate.

Results from the current study suggest that the longevity of female ling is greater than that for males in LIN 3, 4, and 7, but that male ling tend to live longer in LIN 5 and 6. Differing levels of $M$ for fish of the same sex, living in such close proximity, would not be expected. Differences in $M$ between sexes of a species are relatively common when there are marked sexual differences in growth rate, and where there is an apparent difference in $M$, it appears almost universal that the value for females will be lower than that for males, i.e., the longevity of females is greater (Pauly 1980). Ling
exhibit significant differences in growth rates between sexes (Horn 1993b). Comparisons of $Z$ between sexes from individual data sets from LIN 3, 4, and 7 suggest that for males it is about 0.01 to 0.05 higher than for females.

Previous estimates of $M$ for ling were calculated by Horn (1993a), using the $\mathrm{A}_{\max }$ estimator, and slopes of the right hand limb of catch curves from five samples of otoliths collected from 1989 to 1992. (These otolith samples were incorporated in the current study.) Estimates of $Z$ from catch curve regressions ranged from 0.18 to 0.27 , but these were not R1 estimates as they included all non-zero age classes (i.e., they were the RG estimator of Dunn et al. 1999). The $\mathrm{A}_{\max }$ estimates used a $p$ of 0.01 and did not correct for the age classes that are not fully recruited into the sampled fishery. An estimate of $M$ of 0.18 for both sexes was suggested, and has been used in all ling stock assessments since then.

There is little in the current study that can be used to confidently update the estimate of $M$ for ling. No otolith samples are available from an unexploited stock. The estimates that are available vary widely depending on the chosen age at recruitment, and appear to be confounded either by changes in $M$ or $F$ over time, or by uncertainties about the selectivities of some age classes. The CR estimates from the Chatham Rise trawl survey samples (calculated assuming an age at recruitment between 6 and 13 years) suggest a likely $Z$ for females of 0.18 . The assessment of the LIN 3 and 4 stock indicated a virgin biomass of about $150,000 \mathrm{t}$ on the Chatham Rise, and a mean $F$ over the period 1975-92 of 0.02 (Horn 1997). Subtracting the estimated value of $F$ from the $Z$ value for females, suggests that 0.16 is the best value for $M$. Given the conclusion above that $M$ for male ling should probably be slightly higher than that for females, a value of 0.18 is suggested here.

Table B1: Details of samples of catch-at-age data, by Fishstock and sampling period, and estimates of $Z$ from these samples. Method: RT, research trawl; CT, commercial trawl; LL, commercial longline. $N$, number of samples. $R$, age at recruitment. For the CR and R1 estimators, $95 \%$ confidence intervals are plus or minus the value in brackets. For the $A_{\text {max }}$ estimator, two values of $\boldsymbol{p}(\mathbf{0 . 0 1}, 0.05)$ were used. - , not calculated

| Fishstock | Method | Period | $N$ | $\boldsymbol{R}$ | $Z$ (male) |  |  |  |  |  | $Z$ (female) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | CR |  | R1 |  | $\mathbf{A}_{\text {max }}$ |  | CR |  | R1 |  | $\mathrm{A}_{\text {max }}$ |  |
|  |  |  |  |  |  |  |  |  | 0.01 | 0.05 |  |  |  |  | 0.01 | 0.05 |
| LIN 3 \& 4 | RT | 1989-93 | 3 | 12 | 0.26 | (0.02) | 0.26 | (0.04) | 0.27 | 0.27 | 0.21 | (0.01) | 0.22 | (0.03) | 0.27 | 0.23 |
| (Chatham Rise) | RT | 1994-99 | 6 | 12 | 0.19 | (0.01) | 0.22 | (0.04) | 0.23 | 0.25 | 0.19 | (0.01) | 0.21 | (0.03) | 0.23 | 0.23 |
|  | LL | 1998 | 1 | 12 | 0.20 | (0.04) | 0.13 | (0.06) | 0.21 | 0.19 | 0.29 | (0.03) | 0.26 | (0.03) | 0.27 | 0.25 |
| LIN 5 \& 6 | RT | 1989-92 | 4 | 15 | 0.20 | (0.01) | 0.21 | (0.10) | 0.29 | 0.21 | 0.30 | (0.02) | 0.34 | (0.06) | 0.35 | 0.30 |
| (Southern Plateau) | RT | 1993-98 | 3 | 15 | 0.29 | (0.02) | 0.22 | (0.08) | 0.38 | 0.30 | 0.35 | (0.02) | 0.32 | (0.07) | 0.35 | 0.33 |
|  | LL | 1998 | 1 | 15 | 0.22 | (0.05) | - | - | 0.31 | 0.25 | 0.29 | (0.03) | 0.29 | (0.08) | 0.29 | 0.25 |
| LIN 7 | CT | 1991 | 1 | 12 | 0.31 | (0.04) | 0.30 | (0.10) | 0.31 | 0.25 | 0.25 | (0.03) | 0.24 | (0.05) | 0.29 | 0.21 |
| (WCSI) | CT | 1997-98 | 2 | 12 | 0.39 | (0.03) | 0.49 | (0.12) | 0.42 | 0.33 | 0.30 | (0.03) | 0.34 | (0.06) | 0.29 | 0.33 |



Figure B1: Estimated catch-at-age, by sex, for samples of ling taken in trawl surveys of the Chatham Rise, from 1989 to 1999. Lines are the least squares linear regressions fitted to the data points represented by shaded squares.


Figure B2: Estimated catch-at-age, by sex, for samples of ling taken by commercial longline vessels on the Chatham Rise and Southern Plateau. Lines are the least squares linear regressions fitted to the data points represented by shaded squares. (The slope of the regression for the Southern Plateau male sample is not the R1 estimate of $Z$ because some age classes in the regressed range have zero fish.)


Figure B3: Estimated catch-at-age, by sex, for a sample of ling taken in trawl surveys of the Southern Plateau, from 1989 to 1998. Lines are the least squares linear regressions fitted to the data points represented by shaded squares.


Figure B4: Estimated catch-at-age, by sex, for a sample of ling taken by commercial trawlers off the west coast of the South Island, from 1991 to 1998. Lines are the least squares linear regressions fitted to the data points represented by shaded squares.


Figure B5: Estimated catch-at-age, by sex, for samples of ling taken in trawl surveys of the Chatham Rise (as shown in Figure B1), but with two least squares regression lines fitted separately to each data set.


Research trawl-Male


Commercial traw - Male


Figure B6: Estimates of $Z$ using the CR (empty symbols) and R1 (filled symbols) estimation methods and a range of ages at recruitment, for several of the data sets and all sampling methods. Different Fishstocks are represented by different shaped symbols: square, LIN 3 and 4; circle, LIN 5 and 6; diamond, LIN 7.



[^0]:    National Institute of Water and Atmospheric Research

